

$$t \mapsto P(X > t+s \mid X > t) = \underline{P(X > s)}$$

Se $X \sim \mathcal{E}(\lambda)$ la funzione

$$t \mapsto P(X > t+s \mid X > t)$$

non dipende da t (costante in t)

Se c'è un'altra, devo ricordarmi

$$\text{che } t \mapsto P(X > t+s \mid X > t)$$

ha decrescente

$$\begin{aligned}
 P(X > t+s | X > t) &= \\
 &= \frac{P(X > t+s, X > t)}{P(X > t)} = \frac{P(X > t+s)}{P(X > 0)} = \\
 &= \frac{\int_{t+s}^{\infty} \lambda \alpha x^{\alpha-1} e^{-\lambda x^\alpha} dx}{\int_t^{+\infty} \lambda \alpha x^{\alpha-1} e^{-\lambda x^\alpha} dx} = \\
 &= \frac{\int_{(t+s)^\alpha}^{+\infty} \lambda e^{-\lambda y} \underbrace{dy}_{\alpha dy}}{(t+s)^\alpha} = \frac{-e^{-\lambda y} \Big|_{(t+s)^\alpha}^{+\infty}}{\int_t^{+\infty} \lambda e^{-\lambda y} \underbrace{dy}_{\alpha dy} = -e^{-\lambda y} \Big|_t^{+\infty}} = \\
 &= \frac{e^{-\lambda (t+s)^\alpha}}{e^{-\lambda t^\alpha}} = e^{-\lambda (t+s)^\alpha + \lambda t^\alpha}
 \end{aligned}$$

$$t \mapsto e^{-\lambda \left((t+s)^\alpha - t^\alpha \right)} \quad e^- \text{ decrease}$$

$$t \mapsto (t+s)^\alpha - t^\alpha = g(t) \quad e^- \text{ crescente}$$

$$g'(t) = \alpha (t+s)^{\alpha-1} - \alpha t^{\alpha-1} \geq 0$$

$$(t+s)^{\alpha-1} \geq t^{\alpha-1} \implies \alpha-1 \geq 0$$

$$\underline{(\alpha-1) \log(t+s)} \geq \underline{(\alpha-1) \log t}$$

$$\alpha \geq 1$$

$$\alpha < 1$$

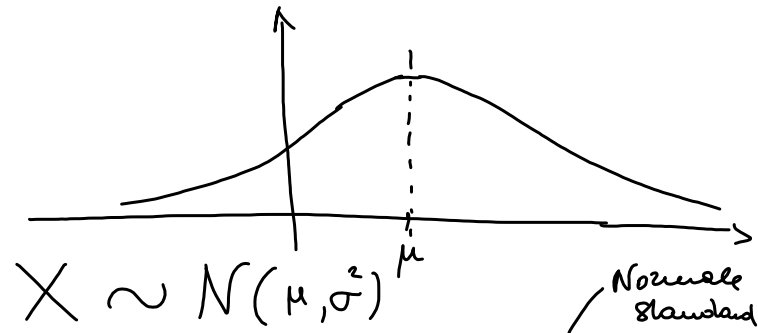
Densità (o legge) Gaussiana
o Normale:

Def. Si dice che X ha legge
gaussiana di parametri μ e σ^2
($\mu \in \mathbb{R}$, $\sigma^2 > 0$) se ha
la seguente densità

$$f(x) = \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, \quad x \in \mathbb{R}$$

$$\int_{-\infty}^{+\infty} f(x) dx = 1$$

$$\int_{-\infty}^{+\infty} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx = \sqrt{2\pi}\sigma$$

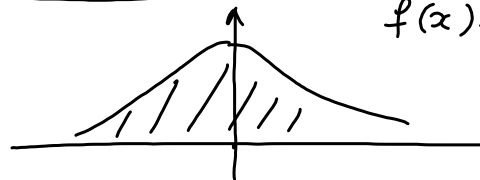


$$X \sim N(\mu, \sigma^2)$$

$$\mu=0, \sigma^2=1 \quad N(0,1)$$

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \quad x \in \mathbb{R}$$

$$f(x) = f(-x)$$



Si dice che X è una v. a.
simmetrica se X e $-X$ hanno
la stessa legge.

Se X ha densità f pari:
($f(x) = f(-x)$), allora X
è simmetrica.

$$\text{Funzione di ripart. di } X = \int_{-\infty}^t f(x) dx$$

$$\text{Funzione di ripart. di } -X$$

$$P(-X \leq t) = P(X \geq -t) =$$

$$= \int_{-t}^{+\infty} f(x) dx = \int_{-t}^{+\infty} f(-x) dx$$

$$\begin{aligned} -x &= y \\ dx &= -dy \end{aligned}$$

$$= - \int_t^{-\infty} f(y) dy = \int_{-\infty}^t f(y) dy$$

$$= P(X \leq t)$$

no!

~~$$X = -X$$~~

Proposizione. Se $X \sim N(m, \sigma^2)$
 Allora la v. a. $Y = aX + b \sim N(am + b, a^2 \sigma^2)$
 $a \neq 0$

Dim.

$$\begin{aligned} P(Y \leq t) &= P(aX + b \leq t) = \\ &= P(aX \leq t - b) = \quad (a > 0) \\ &= P\left(X \leq \frac{t - b}{a}\right) = \\ &= \int_{-\infty}^{\frac{t-b}{a}} f(x) dx = \int_{-\infty}^{\frac{t-b}{a}} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-m)^2}{2\sigma^2}} dx \end{aligned}$$

$$P(Y \leq t) = \int_{-\infty}^{\frac{t-b}{a}} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-m)^2}{2\sigma^2}} dx$$

$$\frac{d}{dt} P(Y \leq t) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{\left(\frac{t-b}{a} - m\right)^2}{2\sigma^2}} \cdot \frac{1}{a}$$

$$= \frac{1}{\sqrt{2\pi}(a\sigma)} e$$

$$= \frac{1}{\sqrt{2\pi}(a\sigma)} e^{-\frac{(t - [b + am])^2}{2a^2\sigma^2}}$$

$$= \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(t - \mu)^2}{2\sigma^2}} \sim N(am + b, a^2\sigma^2)$$

$$X \sim N(m, \sigma^2) \Rightarrow$$

$$Y = aX$$

$$P(Y \leq t) = P(aX \leq t-b) \quad a < 0$$

$$= P\left(X \geq \frac{t-b}{a}\right) =$$

$$= \int_{\frac{t-b}{a}}^{+\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-m)^2}{2\sigma^2}} dx$$

$$= - \int_{+\infty}^{\frac{t-b}{a}} \dots dx$$

$$- \frac{(t - [am+b])^2}{2a^2\sigma^2}$$

$$\frac{d}{dt} P(Y \leq t) = - \frac{1}{\sqrt{2\pi}(a\sigma)} e$$

$$= \frac{1}{\sqrt{2\pi} |a\sigma|} e^{-\frac{(t - [am+b])^2}{2(a\sigma)^2}} \sim N(am+b, (a\sigma)^2)$$

Corollario (i) Se $X \sim N(0,1)$

Allora $Y = \sigma X + \mu \sim N(\mu, \sigma^2)$

(ii) Se $X \sim N(\mu, \sigma^2)$, allora

$$Y = \frac{X - \mu}{\sigma} \sim N(0,1)$$

Dim.

$$(i) Y \sim N(\underbrace{\sigma \cdot 0 + \mu}_m, \underbrace{1 \cdot \sigma^2}_{\sigma^2}) = N(\mu, \sigma^2)$$

$$(ii) Y = \frac{X - \mu}{\sigma} = \underbrace{\frac{1}{\sigma}}_a X + \underbrace{\left(\frac{-\mu}{\sigma}\right)}_b$$

$$Y \sim N(\underbrace{\frac{1}{\sigma} \cdot \mu + \left(\frac{-\mu}{\sigma}\right)}_{= \mu}, \underbrace{\left(\frac{1}{\sigma}\right)^2 \sigma^2}_{= \sigma^2}) = N\left(\frac{1}{\sigma} \cdot \mu + \left(\frac{-\mu}{\sigma}\right), \left(\frac{1}{\sigma}\right)^2 \sigma^2\right) = N(0,1)$$

Se $Y \sim N(0,1)$.

$$\underline{P(Y \leq t)}$$

Se $X \sim N(\mu, \sigma^2)$

$$X \sim N(\mu, \sigma^2)$$

$$Y = \frac{X - \mu}{\sigma} \sim N(0, 1)$$

$$\sigma > 0$$

$$P(\underline{X \leq t}) =$$

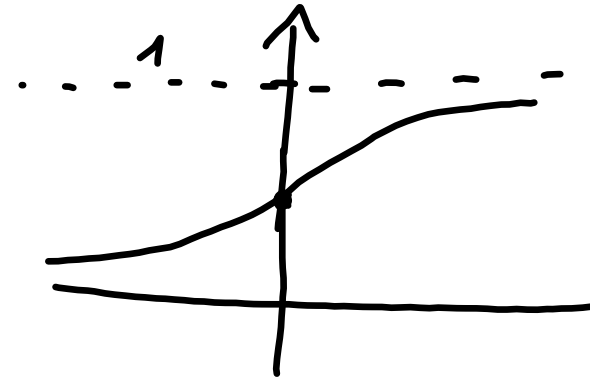
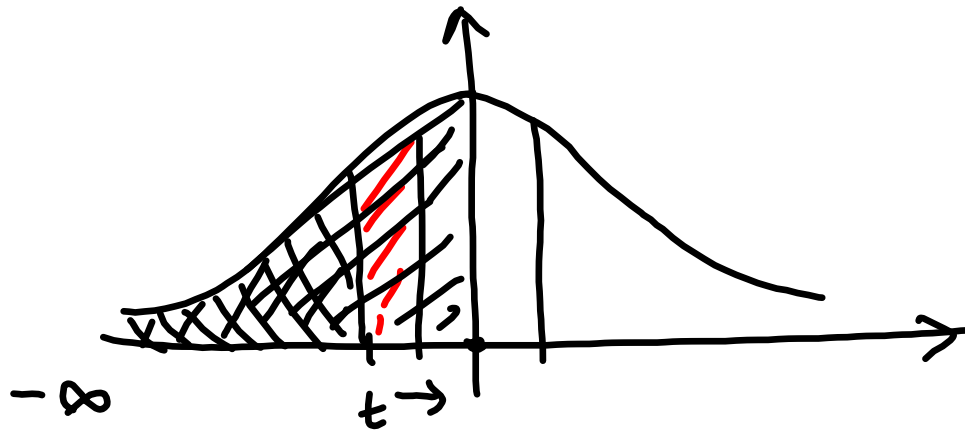
$$= P\left(\frac{X - \mu}{\sigma} \leq \frac{t - \mu}{\sigma}\right) = P\left(Y \leq \frac{t - \mu}{\sigma}\right)$$

Da $Y \sim N(0, 1)$

si pone

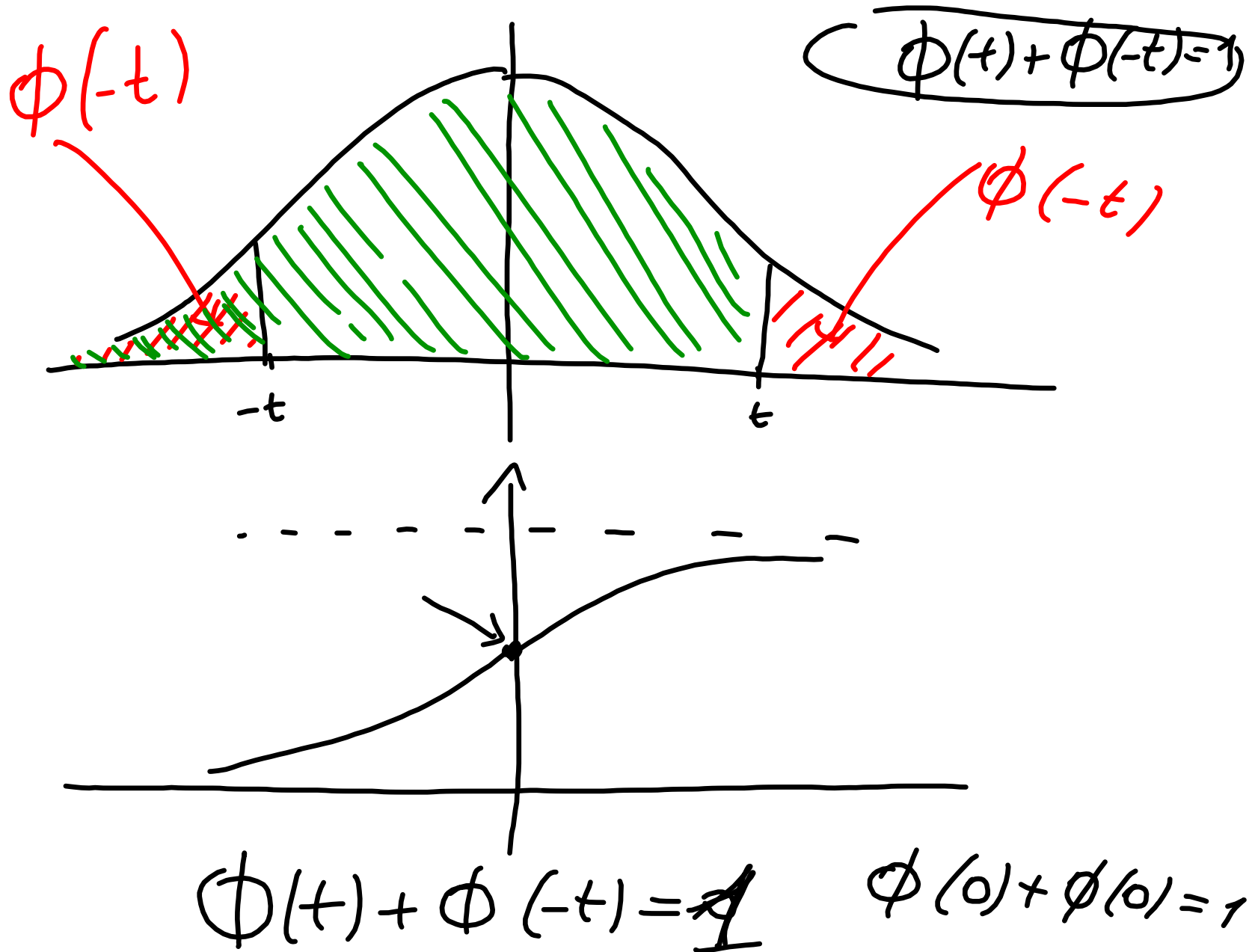
$$\Phi(t) = P(Y \leq t) = \int_{-\infty}^t \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx$$

$$\Phi(t) = \int_{-\infty}^t \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \quad \int e^{-x^2} dx$$



$$\Phi(0) = \frac{1}{2}$$

$$\Phi(t) + \Phi(-t) = 1$$



Sia $X \sim N(\mu, \sigma^2)$. Allora

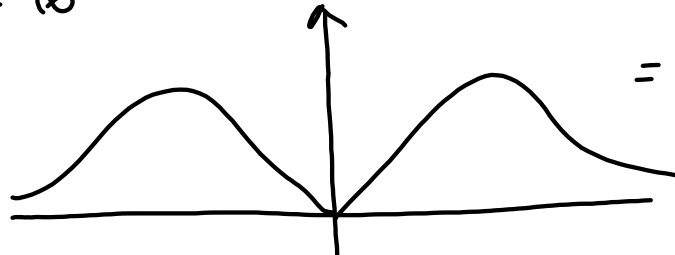
$$E[X] = \mu$$

$$\text{Var} X = \sigma^2$$

Caso della $N(0, 1)$

Sia $X \sim N(0, 1)$

$$E[|X|] = \int_{-\infty}^{+\infty} |x| f(x) dx$$

$$\int_{-\infty}^{+\infty} |x| f(x) dx = \int_{-\infty}^{+\infty} |x| \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$$


$$= 2 \int_0^{+\infty} x \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$$

$$= \frac{2}{\sqrt{2\pi}} \int_0^{+\infty} x e^{-x^2/2} dx = \frac{2}{\sqrt{2\pi}} \left[-e^{-\frac{x^2}{2}} \right]_0^{+\infty}$$

$$= \frac{2}{\sqrt{2\pi}} \cdot 1 = \frac{2}{\sqrt{2\pi}}$$

$$\int x^m e^{-\frac{x^2}{2}} dx$$

$m=0 \int e^{-x^2/2} dx$

$m \geq 1$

se m dispari si

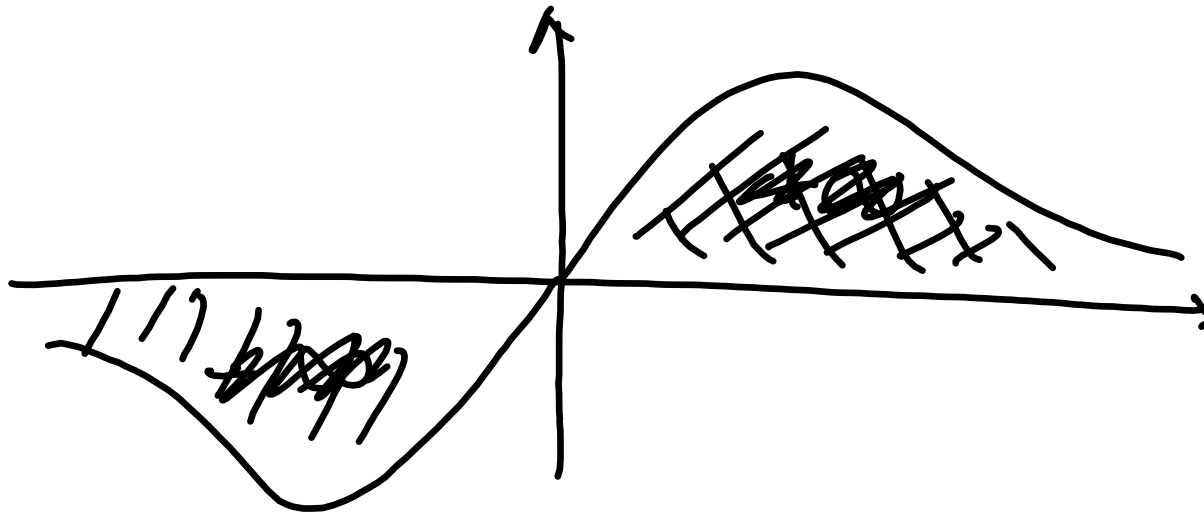
si scrive la primitiva

no se m è pari

$$E[X] = \int_{-\infty}^{+\infty} x f(x) dx =$$

$$= \int_{-\infty}^{+\infty} x \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = 0$$

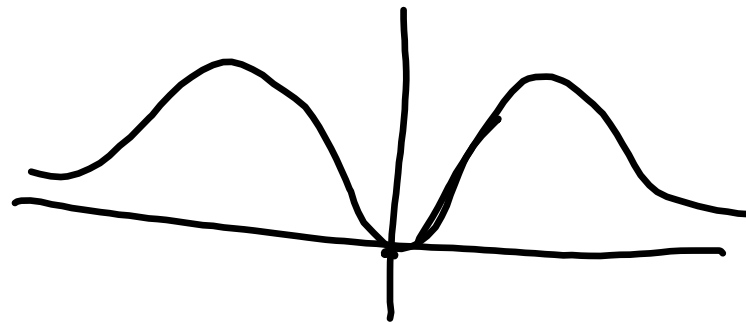
$f(x) = -f(-x)$



$$\text{Var } X = \underbrace{E[X^2]}_{=1} - \underbrace{E[X]^2}_{=0} = 1$$

$$E[X^2] = \int_{-\infty}^{+\infty} x^2 \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$$

$$= \frac{2}{\sqrt{2\pi}} \int_0^{+\infty} x^2 e^{-x^2/2} dx$$



$$\text{Se } X \sim N(0,1), \text{ allora}$$

$$E[X] = 0, \quad \text{Var } X = 1$$

$$\text{Se } X \sim N(\mu, \sigma^2), \text{ allora}$$

$$E[X] = \mu, \quad \text{Var } X = \sigma^2$$

$$Y = \frac{X - \mu}{\sigma} \sim \underline{N(0,1)} \Rightarrow X = \sigma Y + \mu$$

$$E[X] = E[\sigma Y + \mu] = \sigma \underbrace{E[Y]}_{=0} + \mu = \mu$$

$$\text{Var } X = \text{Var}(\sigma Y + \mu) =$$

$$= \text{Var}(\sigma Y) = \sigma^2 \underbrace{\text{Var } Y}_{=1} = \sigma^2$$

Teorema. Dato $X \sim N(\mu_1, \sigma_1^2)$
 $Y \sim N(\mu_2, \sigma_2^2)$ } indipendenti
 Allora la r.a.

$$Z = X + Y \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$$

$$E[Z] = E[X + Y] = E[X] + E[Y] = \mu_1 + \mu_2$$

$$\text{Var } Z = \text{Var}(X + Y) = \text{Var } X + \text{Var } Y = \sigma_1^2 + \sigma_2^2$$

$$X \sim N(m, s^2) \Rightarrow Y = aX + b \sim$$

$$\sim N(am + b, a^2 s^2)$$

$$E[Y] = E[aX + b] = aE[X] + b = am + b$$

$$\text{Var } Y = \text{Var}(aX + b) = a^2 \text{Var } X = a^2 s^2$$

$$\begin{array}{l}
 X \sim N(0,1) \\
 \textcircled{a+b} \quad 2X \sim N(0,4) \\
 X+2X \sim N(0,5) \text{ no!} \\
 3X \sim N(0,9)
 \end{array}
 \left. \begin{array}{l} \\ \\ \\ \end{array} \right\} \text{non sono} \\
 \text{indip.}$$

Più ^{facile} siano X_1, \dots, X_n n v.a. indipend.
 tutte con legge $N(\mu, \sigma^2)$.
 Calc. la legge di

$$\frac{X_1 + \dots + X_n}{n} = \bar{X} = \text{media campione}$$

$$Y = X_1 + \dots + X_n \sim N(n\mu, n\sigma^2)$$

$$\begin{aligned}
 \bar{X} = \frac{X_1 + \dots + X_n}{n} &= \frac{1}{n} Y \sim N\left(\frac{n\mu}{n}, \frac{n\sigma^2}{n^2}\right) \\
 &= N\left(\mu, \frac{\sigma^2}{n}\right)
 \end{aligned}$$

$$\begin{array}{l}
 X \sim N(\mu, \sigma^2) \rightarrow \frac{X - \mu}{\sigma} \sim N(0,1) \\
 \frac{X - E[X]}{\sqrt{\text{Var } X}} = Y
 \end{array}$$

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$$

$$\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} = \sqrt{n} \frac{\bar{X} - \mu}{\sigma} \sim N(0,1)$$

$$= \frac{X_1 + \dots + X_n - n\mu}{\sigma} \sqrt{n} =$$

$$= \frac{X_1 + \dots + X_n - n\mu}{\sigma} \sqrt{n} =$$

$$= \frac{X_1 + \dots + X_n - n\mu}{\sigma \sqrt{n}} = \sqrt{n} \frac{\bar{X} - \mu}{\sigma}$$

$$\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} x^2 e^{-\frac{x^2}{2}} dx = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} x \left(x e^{-\frac{x^2}{2}} \right) dx$$

$$= \frac{1}{\sqrt{2\pi}} \left(- \left. e^{-\frac{x^2}{2}} \cdot x \right|_{-\infty}^{+\infty} - \int_{-\infty}^{+\infty} - e^{-\frac{x^2}{2}} dx \right)$$

$$= \frac{1}{\sqrt{2\pi}} \left(0 + \int_{-\infty}^{+\infty} e^{-\frac{x^2}{2}} dx \right)$$

$$= \int_{-\infty}^{+\infty} \underbrace{\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}}_{=1} dx = 1$$